

JMC

John's Modeling Cult

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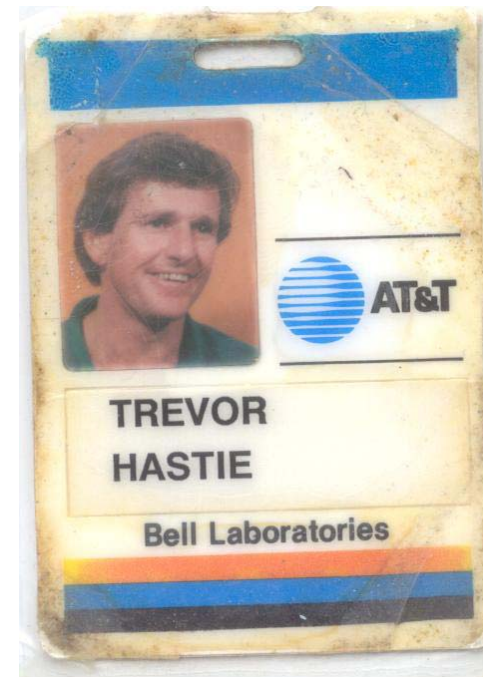
`http://www-stat.stanford.edu/~hastie`



Where and When?

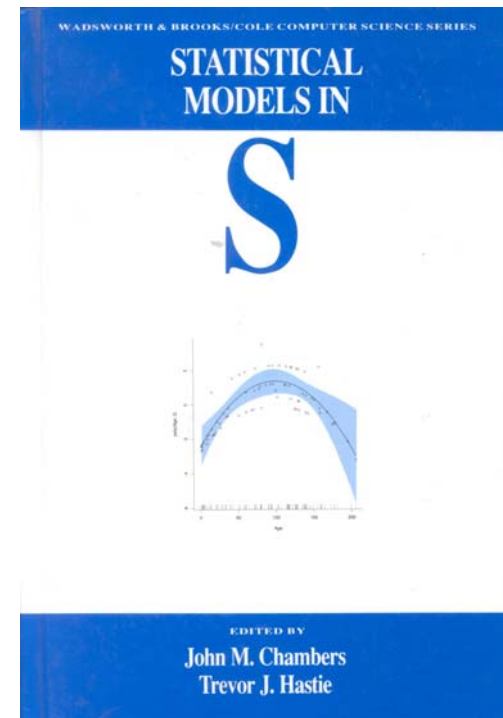
Bell Labs — 1986

- Mid 1985, John gets a phone call from a pushy South African, saying he would rather work at Bell Labs than UBC.
- Eager beaver Hastie joins dept in 1986 - within two weeks visited by Yehuda and Vijay, and told to slow down.
- Very exciting times for a young researcher, with JMC an ideal mentor and role model.



The White Book

- Just finished the GAM book — exhausted — but could not resist being a co-editor of this exciting project with JMC.
- That was before I'd heard of “code review”!
- Just when Daryl and I thought we'd nailed the formula parser, JMC produced his “theorem” (Sec 2.4.1) and cleaned up.
- I am very proud of this collaboration and project.



Rest of the Talk: Regularization Paths

Some of the things I have been doing since leaving the fold.

- Least Angle Regression and the Lasso (with Brad Efron, Iain Johnstone and Rob Tibshirani).

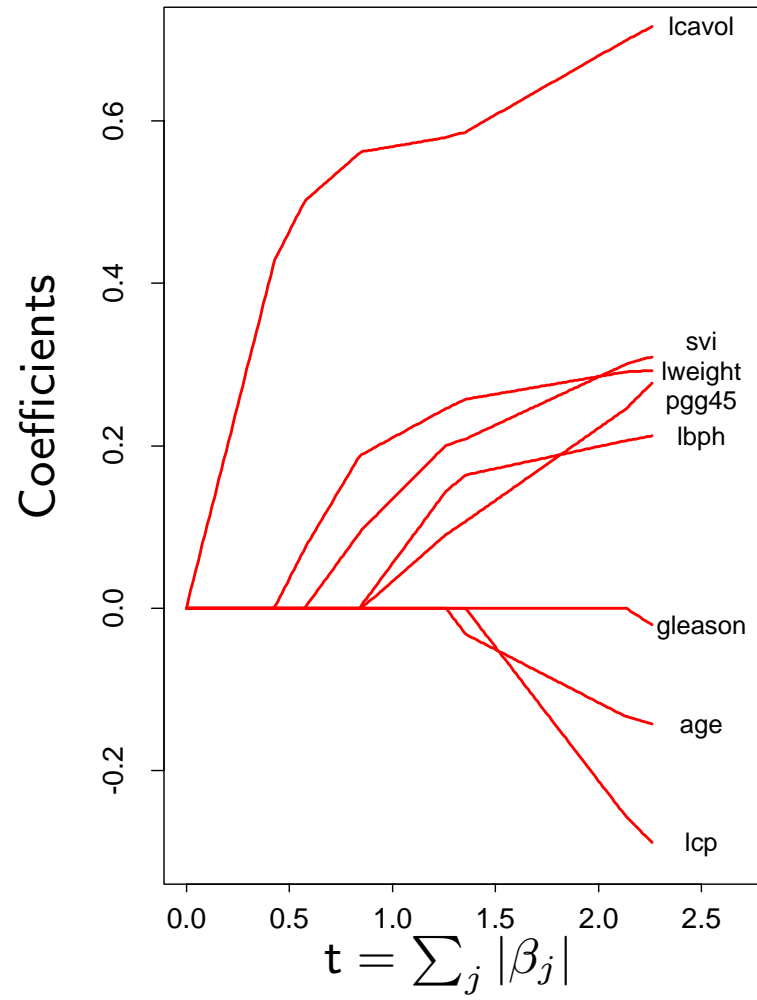


- The SVM path (with Saharon Rosset, Ji Zhu and Rob Tibshirani).

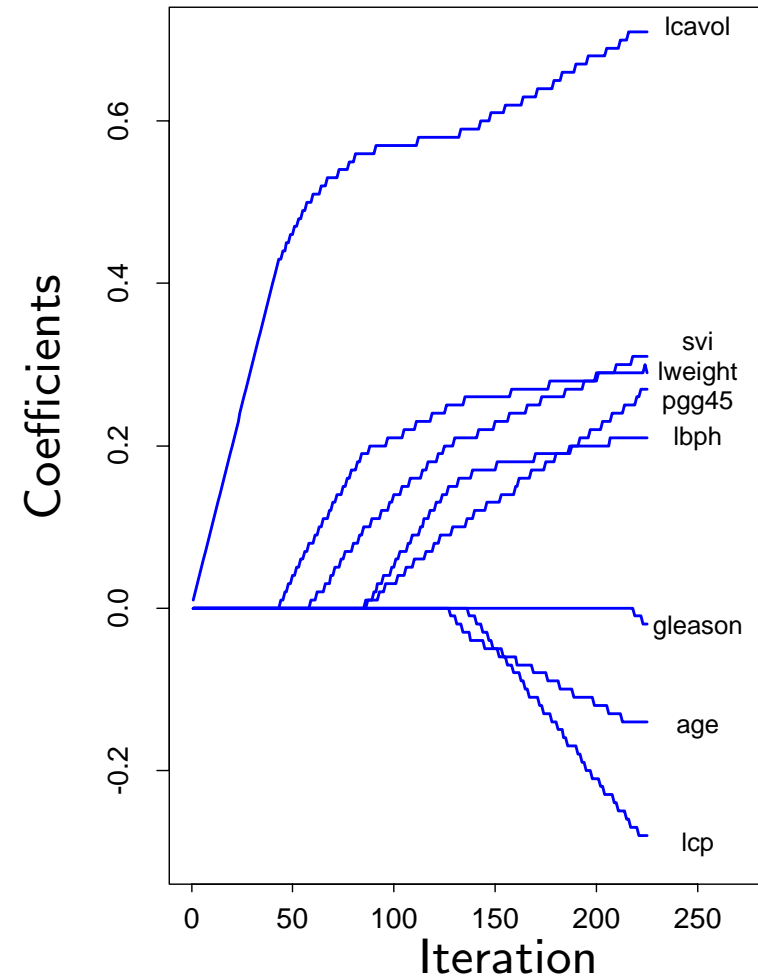


Lasso and Boosting

Lasso



Forward Stagewise



Boosting Linear Regression

Here is a version of least squares boosting for multiple linear regression: (assume predictors are standardized)

(Incremental) Forward Stagewise

1. Start with $r = y$, $\beta_1, \beta_2, \dots, \beta_p = 0$.
2. Find the predictor x_j most correlated with r
3. Update $\beta_j \leftarrow \beta_j + \delta_j$, where $\delta_j = \epsilon \cdot \text{sign}\langle r, x_j \rangle$
4. Set $r \leftarrow r - \delta_j \cdot x_j$ and repeat steps 2 and 3 many times

$\delta_j = \langle r, x_j \rangle$ gives usual forward stagewise; different from forward stepwise

Analogous to regression boosting, with *trees=predictors*

Lasso (Tibshirani, 1995)

- Assume $\bar{y} = 0$, $\bar{x}_j = 0$, $\text{Var}(x_j) = 1$ for all j .
- Minimize $\sum_i (y_i - \sum_j x_{ij}\beta_j)^2$ subject to $\sum_j |\beta_j| \leq s$
- With orthogonal predictors, solutions are soft thresholded version of least squares coefficients:

$$\text{sign}(\hat{\beta}_j)(|\hat{\beta}_j| - \gamma)_+$$

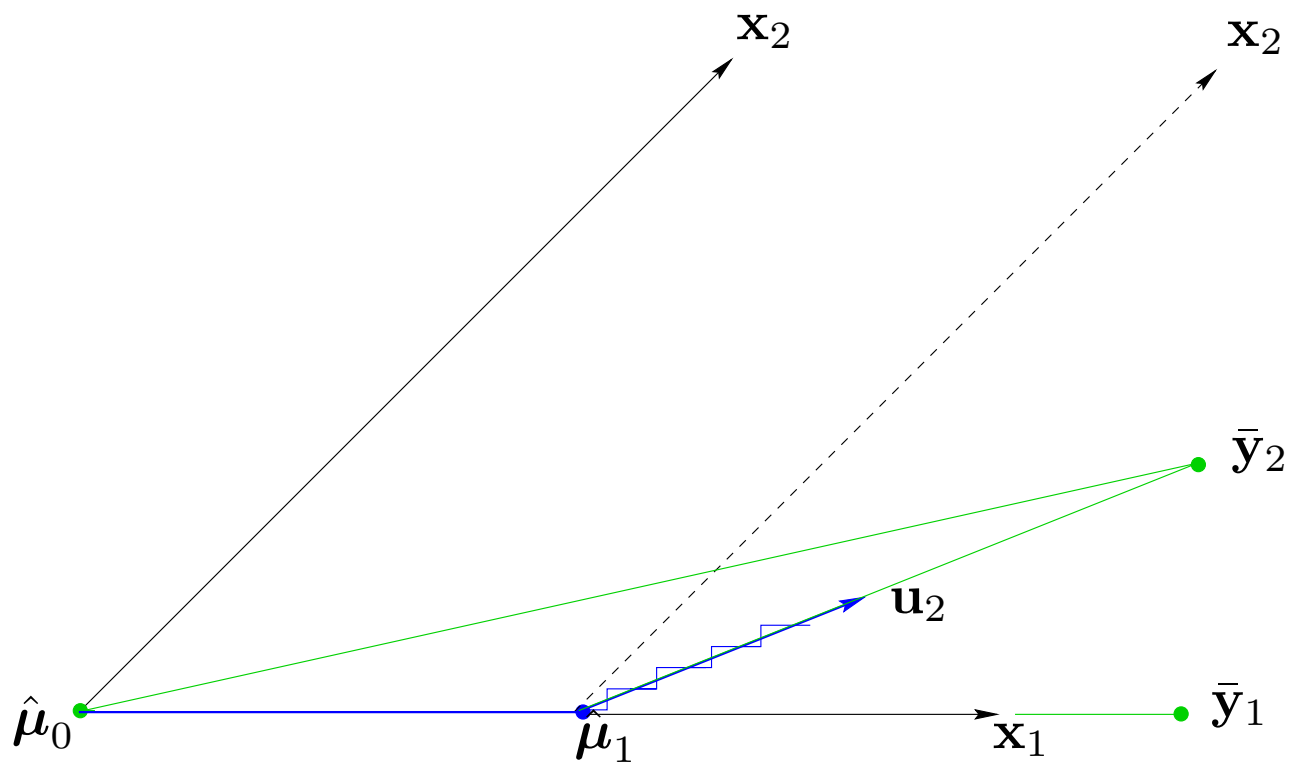
(γ is a function of s)

- For small values of the bound s , Lasso does variable selection.
See pictures

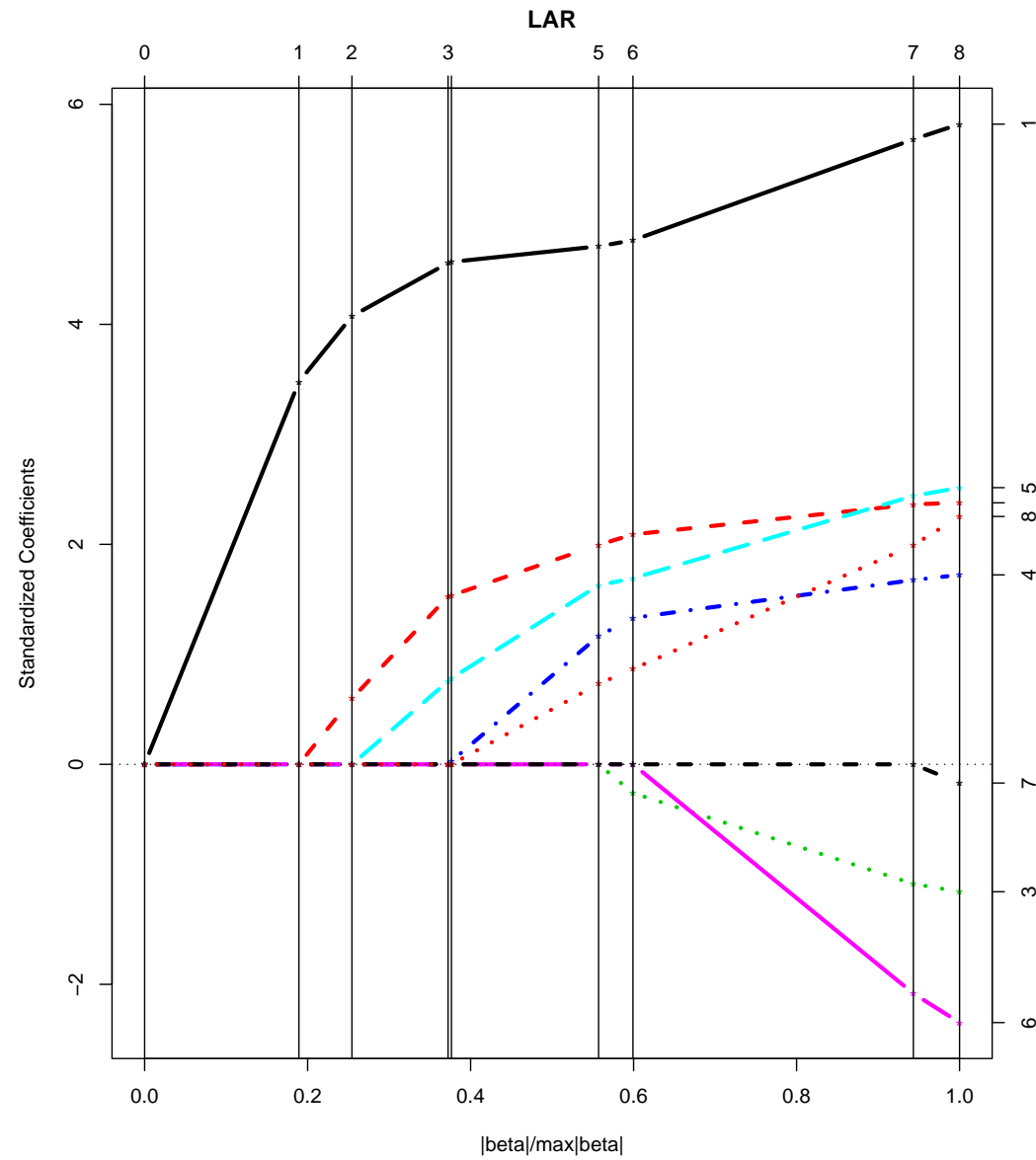
Least Angle Regression — LAR

Like a “more democratic” version of forward stepwise regression.

1. Start with $r = y$, $\hat{\beta}_1, \hat{\beta}_2, \dots, \hat{\beta}_p = 0$. Assume x_j standardized.
2. Find predictor x_j most correlated with r .
3. Increase β_j in the direction of $\text{sign}(\text{corr}(r, x_j))$ until some other competitor x_k has as much correlation with current residual as does x_j .
4. Move $(\hat{\beta}_j, \hat{\beta}_k)$ in the joint least squares direction for (x_j, x_k) until some other competitor x_ℓ has as much correlation with the current residual
5. Continue in this way until all predictors have been entered. Stop when $\text{corr}(r, x_j) = 0 \forall j$, i.e. OLS solution.



The LAR direction \mathbf{u}_2 at step 2 makes an equal angle with \mathbf{x}_1 and \mathbf{x}_2 .



LAR vs Lasso

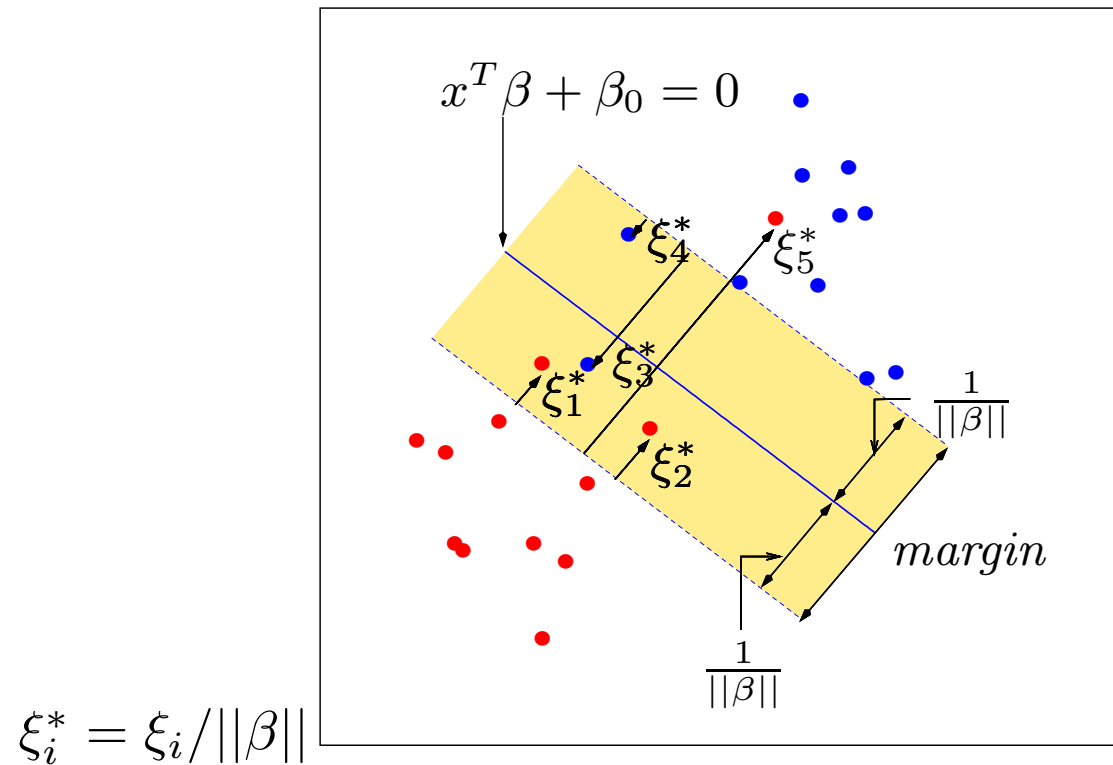
- A modification of LAR fits the entire Lasso path.
- Start with LAR. If a coefficient crosses zero, stop. Drop that predictor, recompute the best direction and continue. This gives the Lasso path
- Proof (lengthy): use Karush-Kuhn-Tucker theory of convex optimization. Informally:

$$\begin{aligned} \frac{\partial}{\partial \beta_j} \{ \|\mathbf{y} - \mathbf{X}\beta\|^2 + \lambda \sum_j |\beta_j| \} &= 0 \\ \Leftrightarrow \\ \langle \mathbf{x}_j, \mathbf{r} \rangle &= \frac{\lambda}{2} \text{sign}(\hat{\beta}_j) \quad \text{if } \hat{\beta}_j \neq 0 \text{ (active)} \end{aligned}$$

Benefits

- Possible explanation of the benefit of “slow learning” in boosting: it is approximately fitting via an L_1 (lasso) penalty
- new algorithm computes entire Lasso path in same order of computation as one full least squares fit. Splus/R Software on my website or CRAN.
- Degrees of freedom formula for LAR:
After k steps, degrees of freedom of fit = k (with some regularity conditions)
- For Lasso, the procedure often takes $> p$ steps, since predictors can drop out. Corresponding formula (conjecture):
Degrees of freedom for last model in sequence with k predictors is equal to k .

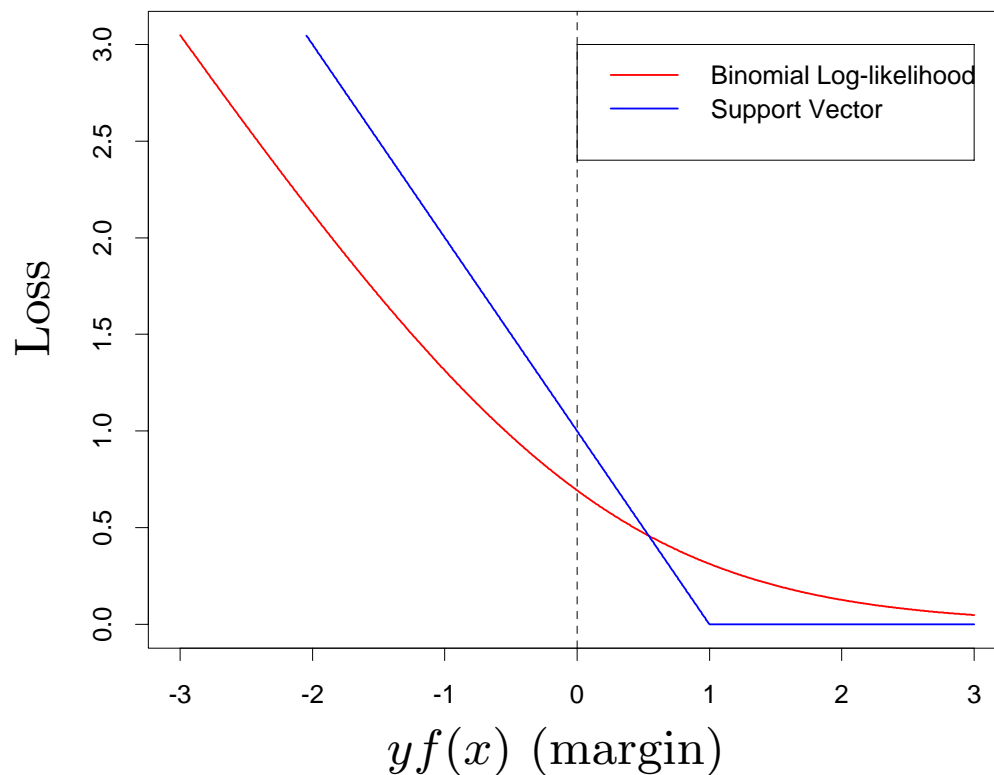
Maximal (Soft) Margin Classifier



$$\min_{\beta, \beta_0} \|\beta\|^2$$

subject to $y_i(x_i^T \beta + \beta_0) \geq 1 - \xi_i$, $\xi_i \geq 0$, $\sum_i \xi_i \leq B$ (*Budget*)

SVM via Loss + Penalty



With $f(x) = x^T \beta + \beta_0$ and $y_i \in \{-1, 1\}$, consider

$$\min_{\beta_0, \beta} \sum_{i=1}^N [1 - y_i f(x_i)]_+ + \frac{\lambda}{2} \|\beta\|^2$$

This *hinge loss* criterion is equivalent to the SVM, with $\lambda \sim B$.

Quadratic Programming

$$L_P : \sum_{i=1}^N \xi_i + \frac{\lambda}{2} \beta^T \beta + \sum_{i=1}^N \alpha_i (1 - y_i f(x_i) - \xi_i) - \sum_{i=1}^N \gamma_i \xi_i$$

$$\frac{\partial}{\partial \beta} : \quad \beta = \frac{1}{\lambda} \sum_{i=1}^N \alpha_i y_i x_i$$

$$\frac{\partial}{\partial \beta_0} : \quad \sum_{i=1}^N y_i \alpha_i = 0,$$

along with the KKT conditions

$$\alpha_i (1 - y_i f(x_i) - \xi_i) = 0$$

$$\gamma_i \xi_i = 0$$

$$1 - \alpha_i - \gamma_i = 0$$

Implications of the KKT conditions

Observations are in one of three states:

$\mathcal{L} = \{i : y_i f(x_i) < 1, \alpha_i = 1\}$, \mathcal{L} for *Left* of the elbow

$\mathcal{E} = \{i : y_i f(x_i) = 1, 0 \leq \alpha_i \leq 1\}$, \mathcal{E} for *Elbow*

$\mathcal{R} = \{i : y_i f(x_i) > 1, \alpha_i = 0\}$, \mathcal{R} for *Right* of the elbow

- Start with λ large, and the margin very wide. All $\alpha_i = 1$. As $\lambda \downarrow 0$, the margin gets narrower.
- For the narrowing margin to pass through a point, its α has to change from 1 to 0 (or from 0 to 1). While this is happening, the point has to *linger* on the margin. Hence the point moves from \mathcal{L} via \mathcal{E} to \mathcal{R} .

The Path

- The α_i are piecewise-linear in λ (or $1/C$) [*MOVIES*].
- The points in \mathcal{E} characterize these paths, since points must stay on the margin ($y_i f(x_i) = 1$) while their α_i lie in $(0, 1)$.
- Points can revisit the margin more than once.
- The coefficients β_0 and β are piecewise-linear in $C = 1/\lambda$.
- The margins can stay wedged while their α_i change, if they are “loaded to capacity”.
- For non-separable data, the loss $\sum_i \xi_i$ achieves a minimum value, with a positive margin.